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LETTER

Operational sex ratio, sexual conflict and the intensity of sexual selection

Patrick S. Fitze^{1,2*} and Jean-François Le Galliard³

¹Museo Nacional de Ciencias Naturales (MNCN-CSIC), Department of Biodiversity and Evolutionary Biology, C/ José Gutiérrez Abascal 2, Madrid 28006, Spain

²Instituto Pirenaico de Ecología (IPE-CSIC), Avenida Regimiento de Galicia s/n, 22700 Jaca (Huesca), Spain

³Laboratoire Fonctionnement et Evolution des Systèmes Ecologiques, CNRS UMR 7625, Université Pierre et Marie Curie, 7 Quai St Bernard, 75005 Paris, France

*Correspondence: E-mail: patrick.fitze@mncn.csic.es

Abstract

Modern sexual selection theory indicates that reproductive costs rather than the operational sex ratio predict the intensity of sexual selection. We investigated sexual selection in the polygynandrous common lizard *Lacerta vivipara*. This species shows male aggression, causing high mating costs for females when adult sex ratios (ASR) are male-biased. We manipulated ASR in 12 experimental populations and quantified the intensity of sexual selection based on the relationship between reproductive success and body size. In sharp contrast to classical sexual selection theory predictions, positive directional sexual selection on male size was stronger and positive directional selection on female size weaker in female-biased populations than in male-biased populations. Thus, consistent with modern theory, directional sexual selection on male size was weaker in populations with higher female mating costs. This suggests that the costs of breeding, but not the operational sex ratio, correctly predicted the strength of sexual selection.

Keywords

Adult sex ratio, *Lacerta vivipara*, operational sex ratio, reproductive costs, sexual conflict, strength of sexual selection.

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INTRODUCTION

Sexual selection results from competition between individuals of one sex for access to gametes of the opposite sex, which leads to a non-random distribution of reproductive success among individuals (Darwin 1871). The investigation of factors affecting the strength of sexual selection is important for understanding sex roles (Andersson 1994), interspecific variation in sexual dimorphism (Møller 1994), and the evolutionary dynamics of sex-specific morphological and behavioural traits (Lande 1980). The most prevalent sexual selection theory predicts that sex roles in reproductive behaviour depend on the operational sex ratio (OSR), defined as the ratio of males to females ready to mate in a population at a given time (Emlen & Oring 1977). This theory is based on the fact that the OSR is a major predictor of the level of mate competition (Kvarnemo & Ahnesjö 1996) and, to some extent, of mate choice (Kokko & Monaghan 2001). The operational sex ratio (OSR) is primarily determined by the adult sex ratio (ASR) and the potential reproductive rate (PRR), i.e. the relative time spent in reproductive vs. non-reproductive activity (Clutton-Brock & Vincent 1991; Parker & Simmons 1996). Therefore, the

intensity of sexual selection should be stronger in the more common sex and the sex with the higher PRR (Bjork & Pitnick 2006); this has been shown for bush crickets (Simmons 1992) and sand gobies (Kvarnemo 1996).

However, this straightforward relationship between the OSR, sex roles, and the intensity of sexual selection is inconsistent with other models of mate competition and mate choice proposing complex interactions between the OSR, parental investment and life history (Kokko & Monaghan 2001; Kokko & Johnstone 2002). Recent evolutionary models suggest that the sign and magnitude of the effect of OSR on the intensity of sexual selection depend on components of the mating system, such as mortality costs (Kokko & Monaghan 2001; Kokko & Johnstone 2002; Simmons & Kvarnemo 2006) – during the time when an individual is not capable of mating (i.e. ‘time out’, Clutton-Brock & Parker 1992) and during the time when an individual is capable of mating (i.e. ‘time in’, Clutton-Brock & Parker 1992) – alternative mating strategies (Mills & Reynolds 2003), the prevalence of coercive mechanisms of sexual selection (Head & Brooks 2006; Simmons & Kvarnemo 2006), or mate encounter rates (Kokko & Monaghan 2001; Kokko & Johnstone 2002). We

investigated the effects of OSR on the strength of the sexual selection in the common lizard (*Lacerta vivipara* Jacquin 1787), whose mating system is characterized by high polygynandry (Fitze *et al.* 2005), significant male aggression during mating (Le Galliard *et al.* 2005), and higher mating costs (mortality during time in) for females than for males when ASRs are biased towards males (Le Galliard *et al.* 2005).

The common lizard breeds annually, has an ovoviviparous reproductive system without parental care, and exhibits a polygynandrous mating system (Laloi *et al.* 2004). Males and females can reach sexual maturity in their first spring and have overlapping home ranges (Le Galliard *et al.* 2005). Adult males emerge from hibernation on average 1 month before the females and become sexually active after their first spring molt (Bauwens *et al.* 1989). Females emerge asynchronously, mate readily after attainment of sexual maturity, and can be receptive during a period of 1–2 weeks (Bauwens & Verheyen 1985). The PRR of males and females is difficult to quantify in many species, including the common lizard (Parker & Simmons 1996). However, experimental manipulation of the ASR within the range of temporal and spatial variability observed in the wild should induce predictable changes in the OSR (Kvarnemo 1996).

We subjected 12 experimental populations – six of which were male-biased and six female-biased – to ASR manipulation and microsatellite-based paternity analysis, including all offspring. Quantification of the strength of sexual selection acting on body size in male and female lizards was based on the relationship between reproductive success (i.e. number of progeny) and body size. Body size, assessed by the lizard's snout-vent length (SVL), is a critical determinant of life history variation in many taxa (Andersson 1994). In female common lizards, body size affects clutch size as well as the timing of reproductive events (Bauwens & Verheyen 1985) and female mate choice (Richard *et al.* 2005). In male lizards, body size affects dominance status and endurance capacity, and therefore the ability to search for females and assure mating (Richard *et al.* 2005). We decomposed reproductive success into four multiplicative fitness components – mating success (number of mates), number of offspring per mate, hatching success and offspring survival until maturity (Clutton-Brock 1988) – and quantified the relationship between these components and body size to estimate the strength of sexual selection acting on body size.

MATERIALS AND METHODS

Experimental manipulation of the ASR

Twelve experimental lizard populations (10 × 10 m) enclosed by plastic walls were established in July 2002 at

the research station at Foljuif (Seine et Marne, France). Each enclosure contained a patch of natural vegetation and two water ponds, which provided natural food and water for the lizards. The enclosures also contained four stone piles for use as basking sites and shelters (for more technical details, see Lecomte & Clobert 1996). Plastic walls prevented lizards from escaping from the enclosures and 18 adult (over 1 year old), 12 yearling (1 year old), and 42–45 newborn juvenile common lizards were released into each of the enclosures at the start of the experiment. In six of the 12 populations, we biased the ASR towards males (MB); in the other six populations the ASR was biased towards females (FB). Fourteen adult males and four adult females were released in MB populations, and four adult males and fourteen adult females in FB populations. The ASR, defined as the number of adult males per number of adult individuals, was 0.78 in MB and 0.22 in FB populations. The sex ratio of the FB populations corresponded to the average ASR of similarly sized patches in a natural population of the native mountainous range in the Cévennes, Southern France (average ASR = 0.22 ± 0.21 SD, 22 patches monitored over 13 years, data provided by M. Massot). The sex ratio of the MB populations corresponded to the largest sex ratios in the same population (see Appendix S1 for further details on the demographic structure of experimental populations).

Assessment of reproductive characteristics

At the end of the mating season, during late May 2003, all live lizards were captured within two consecutive days. The sex ratio of the populations at the end of mating season was still different between treatments (MB populations: 23.7 ± 3.2 SE males, 5.7 ± 0.8 SE females, SR = 0.80 ± 0.03 SE, FB populations: 15.8 ± 2.5 SE males, 20.7 ± 2.1 SE females, SR = 0.40 ± 0.03 SE; $N = 12$, $F_{1,10} = 113.24$, $P < 0.001$). Lizards were not captured during the mating season since this may have disturbed mate choice and mate competition, and thus would have affected the estimates of the intensity of sexual selection. Consequently, the exact numbers of males and females during the mating season is unknown. However, male survival was high in spring and was not affected by ASR manipulation (Le Galliard *et al.* 2005). Thus, the number of males observed corresponds closely to the number of males present during mating in both treatments.

After capture, snout-vent length (SVL) of each lizard was measured to the closest millimetre using a transparent ruler. Captured females were kept in the laboratory in individual terrariums under standardized conditions (water, heat and light regimes standardized according to Le Galliard *et al.*

2003) and were fed with moth larvae every fourth day (*Pyrallis* sp.). After each female laid its eggs, the terrarium was carefully searched for hatchlings and unhatched eggs. Hatching offspring were released after 2 days, into eight initially empty enclosures similar to those used for the ASR experiment. We released the same number of families in each population and made sure that the proportion of families originating from MB and FB populations was the same for each population (see Appendix S2 for detailed population structure and release protocol; Cote *et al.* 2007; Le Galliard *et al.* 2007). All surviving juveniles were recaptured approximately 11 months after hatching, in June 2004. This protocol allowed the longer-term effects of the 2002/2003 ASR on offspring viability selection to be quantified independently of the 2003/2004 ASR.

Genetic protocol and assessment of paternity

We collected a tissue sample of each individual before release, and of each hatchling and unhatched egg in 2003. Females laid a total of 753 eggs, 687 of which hatched. All tissues were stored in 60% ethanol. DNA of all tissue samples was extracted using Perfect gDNA Blood Mini Isolation kit for animal blood (Eppendorf, Hamburg, Germany). We used five to six highly polymorphic microsatellite DNA loci (Lv-3-19, Lv-4-72, Lv-4-alpha, Lv-4-X, Lv-4-115 and Lv-2-145, Boudjemadi *et al.* 1999) to identify fathers. The methods used for DNA extraction, polymerase chain reaction (PCR), and determination of allelic size were previously described (Laloi *et al.* 2004). DNA of all but 12 unhatched eggs could be extracted. Thus, seven unhatched eggs of four different mothers originating from FB populations and five eggs of three different mothers originating from MB populations could not be included in the analyses. We carried out assignment tests using Cervus 2.0 (Marshall *et al.* 1998), to identify the genetic father for each hatchling and unhatched egg (Fitze *et al.* 2005).

Intensity of sexual selection on body size

Some, but not all, common lizards reproduce in their first spring (Boudjemadi *et al.* 1999). Immature female lizards are easily identified by absence of egg production, but it is not easy to distinguish mature males that did not father offspring due to male–male competition or female choice from sexually immature males unable to produce viable sperm. However, male size may indicate attainment of sexual maturity in lizards (Olsson & Madsen 1996). Male size in our populations ranged from 44 to 65 mm ($N = 225$); the smallest male fertilizing eggs had an SVL of 50 mm. There were only 12 potentially immature males (5.3%) with a smaller SVL. Seven males lived in MB and five in FB populations, showing that they were

evenly distributed among enclosures subjected to different ASR manipulations. Thus, we chose a conservative approach and included these males in the sexual selection analysis.

We quantified the strength of sexual selection in male and female lizards based on the relationship between reproductive success (number of progeny surviving until the age of 1 year) and body size (SVL). We decomposed reproductive success W for males and females into four multiplicative fitness components, as follows:

$$W = N_{\text{mate partners}} \times \phi_{N_{\text{eggs/mate partner}}} \times HS \times \Phi_{\text{juvenile}} \quad (1)$$

where $N_{\text{mate partners}}$ corresponds to the number of genetically determined mates, $\phi_{N_{\text{eggs/mate partner}}}$ corresponds to the mean number of eggs per mate, HS to the hatching success (i.e. the probability that an offspring hatched successfully from an egg), and $\Phi_{\text{juvenile survival}}$ is the survival rate from hatching to 1 year of age. Our analysis thus incorporates effects of body size on pre-copulatory sexual selection through mate competition and mate choice, as well as post-copulatory sexual selection through sperm choice, selective abortion of embryos or selective investment into offspring, and viability selection through juvenile survival (Andersson 1994; Eberhard 1996). We only included individuals that were alive during the mating season and our analysis of reproductive success was therefore not affected by natural selection through differential mortality of males and females prior to mating.

Variations in age may have a greater effect on mating tactics than body size; in such cases, the potential effect of body size on sexual selection would be confounded by age (Richard *et al.* 2005). We used long-term data to model the functional relationship between SVL and age in the common lizard (see Appendix S3 and S4 for data collection and analyses). On the basis of this analysis (see below), we calculated the standardized selection gradients for each population and each multiplicative fitness component, using a multiple regression with SVL and age as covariates (Lande & Arnold 1983). To elucidate the selection mechanisms, we investigated whether the standardized fitness components were correlated, using simple regressions, and estimated the importance of each multiplicative fitness component for reproductive success, using a multivariate regression with the standardized reproductive success (W) as a dependent variable and the standardized fitness components as covariates (Conner *et al.* 1996). We performed one-way ANOVA or Wilcoxon-signed ranks tests to determine whether the standardized selection gradients were significantly different between sex ratio manipulations. The quadratic selection gradients testing for stabilizing or disruptive selection were not significant. Hence, only the results of linear selection gradients (β) are reported here.

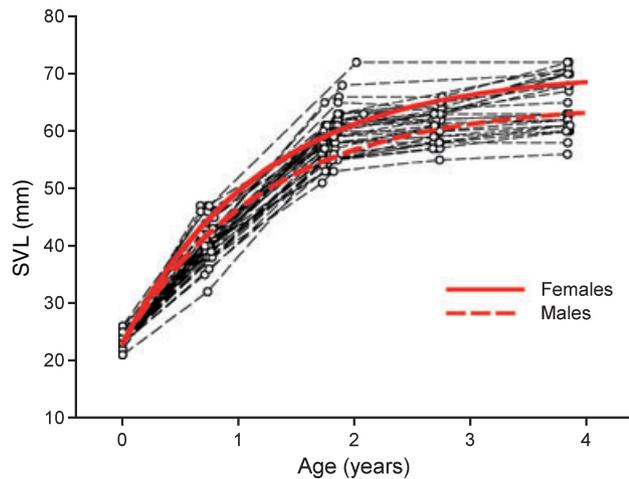


Figure 1 Body size growth trajectories of male and female common lizards. Body size growth trajectories of male and female common lizards from, birth to the age of 4 years, measured in experimental populations similar to, and from the same meadow as, those used for the ASR experiment (see Appendix S3 for further details). Individual growth curves for 42 lizards monitored until the age of 3 or 4 years (dashed black lines) are shown, with sex-specific estimates of the best fitting nonlinear growth model (red lines). The van Bertalanffy growth model included individual variation in asymptotic body size ($\chi^2 = 10.26$, $P = 0.001$) and exponential growth rate ($\chi^2 = 6.33$, $P = 0.01$), as well as sex differences in asymptotic body size ($F_{1,167} = 28.54$, $P < 0.0001$). The predicted asymptotic body size was 70.3 mm (68.5, 72.1) 95% CI for females and 64.8 mm (63.2, 66.2) for males. See Appendix S3 and S4 for details.

RESULTS

The relationship between body size and age – ranging from 1 to 4 years – for common lizards was best described by a nonlinear growth model (Fig. 1). Females grew to a larger asymptotic body size than males (Fig. 1). Although common lizards show indeterminate growth, female and male lizards reached 86% and 85% of their asymptotic body size by the age of 2 years. Body size ranks of individual lizards remained constant after 2 years of age ($\chi^2 = 42.71$, $P < 0.001$, repeatability of individual body size: $r = 0.65$), but not from 1 year of age ($\chi^2 < 0.001$, $P > 0.99$, $r < 0.001$); this suggests the existence of compensatory growth early in life. The combination of fast growth early in life with significant and predictable individual variation in body size among adults enabled us to assess the combined effects of age and size on sexual selection without the complication of strong colinearity between age and body size (Quinn & Keough 2002).

Current theory suggests that the strength of sexual selection in each sex is accurately predicted by the OSR, which is determined by the ASR (Emlen & Oring 1977). It is

generally assumed that the limiting sex is more selective in mate choice, thereby determining the strength of sexual selection in the predominant sex, which invests more in competition (Kvarnemo & Ahnesjö 1996). Thus, sexual selection should be stronger for males in MB than in FB populations, whereas sexual selection in females should be stronger in FB than in MB populations. The number of mates was under strongest selection in our study; this is consistent with Bateman's principle of stronger selection for access to mates than for other components of sexual selection (Shuster & Wade 2003; Fig. 2). On average, there was a positive correlation between the number of mates and body size in both sexes, indicating positive, directional sexual selection acting on body size. The sign of the effect of ASR manipulation on this gradient of sexual selection clearly opposes predictions based on the prevalent sexual selection theory (Emlen & Oring 1977). The intensity of sexual selection acting on male body size, measured by the number of mates, was 4.4 times higher in populations where males were rare (FB: $\beta = 0.518 \pm 0.03$ SE) than in populations where males were abundant (MB: $\beta = 0.117 \pm 0.04$ SE; $F_{1,10} = 9.614$, $P = 0.011$, Fig. 2a). As for males, female body size was under stronger selection when females were rare. The intensity of sexual selection on female body size, as assessed from the number of male partners, was three times higher in MB ($\beta = 1.345 \pm 0.14$ SE) compared with FB populations ($\beta = 0.454 \pm 0.04$ SE; Wilcoxon $W(WW) = 0$, $P = 0.008$, Fig. 2b). Age was not a good predictor of sexual selection (all $P > 0.1$). Furthermore, there were no significant effects of ASR manipulation on the correlations between age (all $P > 0.05$) or body size and the number of eggs per mate (males: $\beta_{FB} = 0.017 \pm 0.08$ SE; $\beta_{MB} = 0.431 \pm 0.17$ SE, $F_{1,10} = 0.815$, $P = 0.388$; females: $\beta_{FB} = 0.496 \pm 0.08$ SE; $\beta_{MB} = -0.067 \pm 0.14$ SE, $F_{1,8} = 3.10$, $P = 0.116$), hatching success (males: $\beta_{FB} = 0.397 \pm 0.10$ SE; $\beta_{MB} = 0.291 \pm 0.07$ SE, $F_{1,9} = 0.126$, $P = 0.731$; females: $\beta_{FB} = 0.132 \pm 0.12$ SE; $\beta_{MB} = 2.099 \pm 0.53$ SE, $WW = 5$, $P = 0.166$), or juvenile survival (males: $\beta_{FB} = 0.156 \pm 0.13$ SE; $\beta_{MB} = 0.128 \pm 0.596$ SE, $F_{1,7} = 0.001$, $P = 0.973$; females: $\beta_{FB} = -0.041 \pm 0.10$ SE; $\beta_{MB} = -6.768 \pm 2.66$ SE, $WW = 12$, $P = 0.067$, Fig. 2).

DISCUSSION

Until recently, it has been widely accepted that the strength of sexual selection is mainly determined by the operational sex ratio and that sexual selection is stronger in the abundant sex (Emlen & Oring 1977). However, the findings from a number of empirical studies are not consistent with these predictions. For example, Van Dongen *et al.* (1998) found that male, but not female, winter moths (*Operophtera brumata*) were choosing mates, despite a heavily male-biased

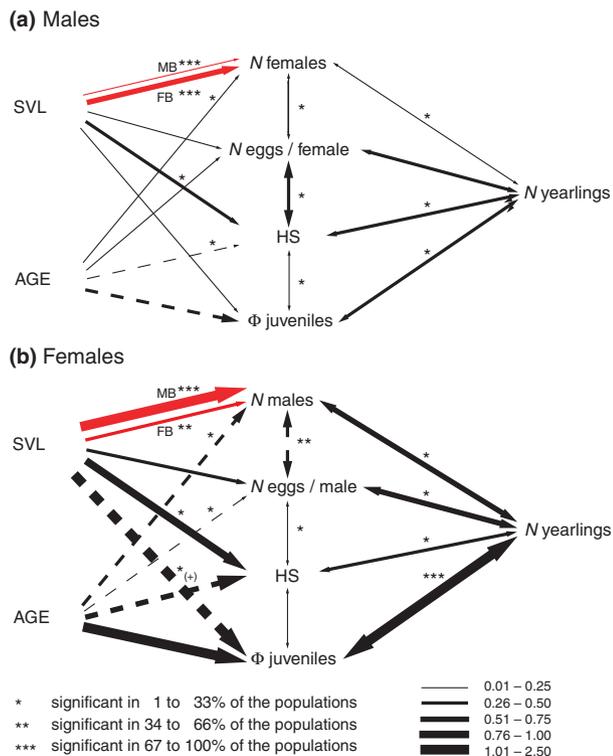


Figure 2 Path diagrams of selection acting on fitness. Path diagrams for selection acting on the four multiplicative fitness components of selection in male (a) and female (b) common lizards. Each arrow represents the mean standardized selection gradient for the twelve experimental populations (six MB and six FB populations). Significant differences in the strength of selection between MB and FB populations are represented by separate arrows. Correlations are depicted as double-headed arrows, and causal relationships as single-headed arrows. Dashed arrows denote negative coefficients. Arrow width is proportional to the standardized coefficient (see scale). Stars indicate the significance level of the respective relationships. HS refers to hatching success and Φ refers to juvenile survival (see eqn 1). Red arrows depict significant differences in the strength of selection between female-biased (FB) and male-biased (MB) populations (for further details see Appendix S5). For males, the intensity of sexual selection on the number of mates was significantly higher in FB than in MB populations ($F_{1,10} = 9.614$, $P = 0.011$); for females, the intensity of sexual selection on the number of mates was significantly higher in MB populations ($WW = 0$, $P = 0.008$). For males in FB populations, reproductive success was mainly determined by the number of mates ($R^2 = 0.206 \pm 0.100$ SE), whereas the number of eggs fertilized per female ($R^2 = 0.061 \pm 0.004$ SE), the hatching success ($R^2 = 0.094 \pm 0.083$ SE), and the juvenile survival ($R^2 = 0.053 \pm 0.019$ SE) accounted for less of the variance in total fitness. For females in MB populations, reproductive success was mainly determined by juvenile survival ($R^2 = 0.491 \pm 0.117$ SE), and to a lesser extent by the number of mates ($R^2 = 0.065 \pm 0.021$ SE), the number of eggs per mate ($R^2 = 0.044 \pm 0.014$ SE), and hatching success ($R^2 = 0.077 \pm 0.066$ SE).

OSR. Recently, Kokko and Monaghan (2001) and Kokko and Johnstone (2002) designed an evolutionary model showing that the OSR may be a good predictor of the intensity of sexual selection in some circumstances, but that parental investment and mating costs are the primary determinants of sex roles. The factors promoting 'choosiness' in one sex are high sex-specific mortality costs of breeding (mortality during 'time out' and relative duration of 'time out'), high sex-specific mate-encounter rate, low sex-specific mortality during 'time in', and highly variable quality of the opposite sex (Kokko & Johnstone 2002).

In this study, we manipulated the OSR in replicated populations and tested the effect of MB OSR on the strength of sexual selection acting on body size. We assessed reproductive success (the number of genetic offspring) in adults surviving until the mating season and quantified the intensity of sexual selection with standardized selection gradients – a common practice for measuring the intensity of sexual selection (Shuster & Wade 2003). This approach excludes differential survival selection between individuals subjected to different ASR treatments and thus allows exclusive quantification of sexual selection intensity. Furthermore, we decomposed reproductive success into four multiplicative terms: the number of genetic mates, the number of genetic offspring per mate, the hatching success of the genetic offspring, and the juvenile survival of the genetic hatchlings. Hence, we were able to investigate the effects of OSR variation on pre-copulatory and post-copulatory components of sexual selection in far more detail than in the previous studies.

Our findings are consistent with predictions based on the sexual selection models recently proposed by Kokko and Monaghan (2001) and Kokko and Johnstone (2002). We previously reported that mating costs for females during 'time in' were higher in MB than in FB populations due to sexual harassment by males (Le Galliard *et al.* 2005). Further, females of MB populations produced fewer offspring during the ASR manipulation than females of FB populations (Le Galliard *et al.* 2005) and had higher survival rates in the year following the ASR manipulation (Le Galliard *et al.* 2007). Thus, breeding costs of females of MB populations seem to be lower than those of FB populations. Mate encounter rates and variation in mate quality did not differ between enclosures subjected to different ASR manipulations (Fitze *et al.* 2005). Similarly, variation in mortality during 'time out' did not differ between females or males subjected to different ASR perturbations (Le Galliard *et al.* 2005). According to the recent sexual selection models (Kokko & Monaghan 2001; Kokko & Johnstone 2002), both higher costs of mating (higher mortality during 'time in') and lower costs of breeding for females of MB populations should reduce

female choosiness. Consequently, there should be weaker sexual selection – in terms of access to mates – on male body size. Indeed, we found that the strength of sexual selection on male SVL was weaker in MB than in FB populations. The classic models of sexual selection predict that the competition among males for access to females should be higher in MB than in FB populations (Emlen & Oring 1977). However, our findings are clearly inconsistent with these models: the ASR-induced shift in the strength of sexual selection in males was opposite to that predicted by classical models. Furthermore, ASR manipulation had no effect on male survival (Le Galliard *et al.* 2005) and on survival selection acting on male body size between July 2002 and May 2003 (see Appendix S6); this indicates that the strength of intrasexual competition between males was similar between enclosures subjected to different sex ratio manipulations.

The recent sexual selection models also predict that increased selection in one sex may decrease selection in the other sex (Kokko & Monaghan 2001; Kokko & Johnstone 2002). Again, our experimental results are consistent with this prediction: the intensity of selection acting on female body size, as assessed by the number of mates, was higher in MB populations than in FB populations. Increased selection on female body size in MB populations may potentially result from size-dependent costs of increased male aggression for females of MB populations (Head & Brooks 2006). Larger females from MB but not from FB populations showed more mating scars than small females, and larger females may be less sensitive to increased male aggression than smaller females (Fitze *et al.* 2005; Le Galliard *et al.* 2005). Differences in selection on female body size may also arise due to male mate choice. Yet, given that male mate choice has been rarely reported in reptiles (Orrell & Jenssen 2002), male mate choice is less likely to be responsible for stronger selection on females of MB populations.

The observed pattern summarized for females in Figure 2 may not necessarily imply sexual selection if the number of mate partners is not causally linked with reproductive success (Shuster & Wade 2003). For example, larger females might be more harassed by males and consequently they may have mated more often, but female body size may independently determine clutch size, hence reproductive success. Alternatively, females may benefit from multiple mating through increased ovulation (Fitze *et al.* 2005; Eizaguirre *et al.* 2007) and thus the number of mate partners may be causally linked with reproductive success. We tested this hypothesis by analysing the relationship between clutch size, female size and multiple mating, assessed by the number of mating scars – a good proxy for the number of mating attempts (Fitze *et al.* 2005). We found that clutch size, independent of body size, increased with the number of mating scars, and that the increase did not significantly

differ between ASR treatments, i.e. no significant interaction between ASR and number of mating scars existed [N mating scars: $F_{1,143} = 12.99$, $P < 0.001$, estimate: 0.211 ± 0.06 SE; SVL $F_{1,143} = 240.65$, $P < 0.001$, estimate: 0.204 ± 0.01 SE; ASR effect: $F_{1,10} = 12.62$, $P = 0.005$, estimate(FB): 0.441 ± 0.12 SE; interaction: ASR \times N mating scars $F_{1,142} = 0.105$, $P = 0.746$, estimate: 0.020 ± 0.06 SE]. This suggests that multiple mating, independent of body size, led to increased clutch size, indicating that the relationship between multiple mating and reproductive success could be causal in female common lizards (Fitze *et al.* 2005; Eizaguirre *et al.* 2007) and indicating that sexual selection may also be responsible for the observed patterns in females. Future studies should further clarify the link between multiple mating and female fecundity and the importance of sexual selection through multiple mating in females.

Our study demonstrates a relationship between OSR and the intensity of sexual selection opposite to that predicted by classical sexual selection theory (Emlen & Oring 1977). However, our findings are consistent with recent sexual selection models (Kokko & Monaghan 2001; Kokko & Johnstone 2002) that emphasize the importance of measuring both mating and breeding costs to understand sexual selection. Kokko & Monaghan (2001) suggested that studies of sexual selection in species with flexible mating systems and life histories may be useful for elucidating how OSR and life history predict the intensity of sexual selection. In a recent experiment, Simmons & Kvarnemo (2006) showed that in the Australian tettigoniid, *Kawanaphila narree*, PRR, which were determined by the time required for a breeding attempt, predicted sexual selection better than breeding costs. This demonstrates that, in contrast to the common lizard, variation in mating and breeding mortality (cost of breeding) is not an important determinant of sexual selection in *K. narree*. Kokko & Monaghan (2001) and Kokko & Johnstone (2002) stated that 'if they were to simplify sexual selection to a single parameter that best predicts a mating system', it would be the mortality costs of breeding. However, this simplification would lead to incorrect predictions in both our and Simmons & Kvarnemo's (2006) studies; these studies thus demonstrate that the complexity of sexual selection cannot be reduced to a single parameter.

Our findings have important implications for our understanding of the ecological context of sexual selection and the long-term evolution of sexually dimorphic characters, such as body size. They highlight that the balance between mate choice, mate competition and sexual conflict may determine how biased sex ratios influence the strength of sexual selection on body size. Intersexual conflicts may therefore affect the evolution of sexual size dimorphism (Szekely *et al.* 2004), including growth and reproductive tactics (Ramm 2007), and they may significantly determine

the speciation (Gavrilets & Waxman 2002; Masta & Maddison 2002; Martin & Hosken 2003) of sexually reproducing organisms.

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SUPPLEMENTARY MATERIAL

The following supplementary material is available for this article:

Appendix S1 Detailed demographic structure of experimental populations.

Appendix S2 Demographic structure of experimental populations during the second year of the study.

Appendix S3 Long-term study of growth patterns.

Appendix S4 Statistical analysis of growth trajectories.

Appendix S5 Statistical analysis of the intensity of sexual selection.

Appendix S6 Statistical analysis of survival selection acting on male body size and body condition.

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